

Error Reduction from Stacked Regressions

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Joint work with Xin Chen (Princeton) & Yan Shuo Tan (NUS)

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• Method has found widespread applications (finance, healthcare, commerce, ..., Kaggle competitions)

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- Better to make weights non-negative, i.e., $\alpha_k \ge 0$
- Minimizers $\hat{\alpha}_k$ yield stacked model $\hat{\mu}_{\text{stack}}(x) = \sum_{k=1}^{M} \hat{\alpha}_k \hat{\mu}_k(x)$

This resulting predictor $\sum_{k=1}^{M} \hat{\alpha}_k \hat{\mu}_k(x)$ appears to almost always have lower prediction error than the single prediction $\hat{\mu}_k$ having lowest cross-validation error. The word "appears" is used because a general proof is not yet in place. — Leo Breiman (1996) This resulting predictor $\sum_{k=1}^{M} \hat{\alpha}_k \hat{\mu}_k(x)$ appears to almost always have lower prediction error than the single prediction $\hat{\mu}_k$ having lowest cross-validation error. The word "appears" is used because a general proof is not yet in place. — Leo Breiman (1996)

Goal of talk is to theoretically confirm this in certain cases

- M = 50 nested regression trees $\hat{\mu}_k$ from pruning
- Weights $\hat{\alpha}_k$ sum to 0.96

Data Set				
	Housing		Ozone	
	Best	Stacked	Best	Stacked
Error	20.9	19.0	23.9	21.6

Table 1 Test Set Prediction Errors

Table 2. Stacking Weights

# Terminal Nodes	Weight	
7	.29	
10	.13	
23	.13	
26	.09	
29	.12	
34	.20	

- M = 40 linear models $\hat{\mu}_k$ from stepwise deletion
- Weights $\hat{\alpha}_k$ sum to between 0.7 and 0.9



• Nonparametric regression with fixed design and known variance σ^2 :

$$y_i = \mu(x_i) + \sigma \varepsilon_i, \quad i = 1, 2, ..., n, \quad \varepsilon_i \stackrel{\text{iid}}{\sim} N(0, 1)$$

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- Training error of $\hat{\mu}$ is

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- Define best single model $\hat{\mu}_{\text{best}}$ as $\hat{\mu}_{\hat{k}}$, where

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- Will describe performance of $\hat{\mu}_{\text{stack}}$ relative to $\hat{\mu}_{\text{best}}$

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- In both cases, estimators µ̂_k are least-squares projections of *y* onto nested subspaces A₁ ⊂ A₂ ⊂ · · · ⊂ A_M
- · Because models are nested,

$$d_1 < d_2 < \cdots < d_M$$

and

$$\|\mathbf{y} - \hat{\mu}_M\|^2 < \cdots < \|\mathbf{y} - \hat{\mu}_2\|^2 < \|\mathbf{y} - \hat{\mu}_1\|^2$$

· Ideally want weights to minimize expected in-sample error

$$\mathsf{Err}(\boldsymbol{\alpha}) = \mathbb{E}\bigg[\left\|\boldsymbol{\mu} - \sum_{k=1}^{M} \alpha_k \hat{\boldsymbol{\mu}}_k\right\|^2\bigg],$$

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• Training error
$$\boldsymbol{R}(\boldsymbol{\alpha}) = \left\| \boldsymbol{y} - \sum_{k=1}^{M} \alpha_k \hat{\mu}_k \right\|^2$$

• Degrees of freedom df(α) = $\sum_{k=1}^{M} \alpha_k d_k$

· Solve quadratic program with linear constraints:

minimize
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• So $R(\hat{\alpha}) + \frac{2\sigma^2}{n} df(\hat{\alpha}) - \sigma^2$ is no longer unbiased estimator of error for stacked model with adaptive weights

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$$\begin{array}{ll} \mbox{minimize} & R(\alpha) + \frac{2\sigma^2}{n} \mbox{df}(\alpha) - \sigma^2 + \frac{\sigma^2}{n} \frac{(\lambda - 1)^2}{\lambda} \mbox{dim}(\alpha) \\ \mbox{subject to} & \alpha_k \geqslant 0, \quad k = 1, 2, \dots, M \end{array}$$

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 - · Same complexity as finding best single model
- Solution satisfies $\sum_{k=1}^{M} \hat{\alpha}_k < 1$, despite no explicit sum constraint

• Due to nested structure and non-negative constraints, problem reduces to weighted isotonic regression:

minimize
$$\sum_{k=1}^{M} w_k (z_k - \gamma_k)^2$$

subject to $\gamma_1 \leqslant \gamma_2 \leqslant \cdots \leqslant \gamma_M$,

where $w_k = \|y - \hat{\mu}_{k-1}\|^2 - \|y - \hat{\mu}_k\|^2$ and $z_k = (d_k - d_{k-1})/w_k$

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Closed-form solution:

$$\hat{\gamma}_{k} = \frac{\sigma^{2}}{n} \min_{k \le i \le M} \max_{0 \le j < k} \frac{d_{i} - d_{j}}{\|y - \hat{\mu}_{j}\|^{2} - \|y - \hat{\mu}_{i}\|^{2}}$$

The following representations hold:

$$\hat{\mu}_{\textit{best}}(\boldsymbol{x}) = \sum_{k=1}^{M} (\hat{\mu}_{k}(\boldsymbol{x}) - \hat{\mu}_{k-1}(\boldsymbol{x})) \mathbf{1}(\hat{\gamma}_{k} < 1/\lambda),$$

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 \u03c6 k_k(x) - \u03c6 k_{k-1}(x)
 across successive sub-models
- Stacked model additionally shrinks these predictive differences towards zero by factor $(1 \hat{\gamma}_k)$
- Performs model selection and adaptive shrinkage simultaneously

Theorem (Chen, K., & Tan, 2023) If $d_k \ge d_{k-1} + 4$ for all k, then $\mathbb{E}[\|\mu - \hat{\mu}_{\text{stack}}\|^2] < \mathbb{E}[\|\mu - \hat{\mu}_{\text{best}}\|^2]$

· Theoretically confirms Breiman's empirical findings

- Error gap $\mathbb{E}\big[\|\mu-\hat{\mu}_{\rm best}\|^2-\|\mu-\hat{\mu}_{\rm stack}\|^2\big]$ can be lower bounded by

$$\frac{\sigma^2}{n} \mathbb{E}\left[\min_{1 \le k \le M} \frac{(d_k - 4k)^2}{(n/\sigma^2)(\|\mathbf{y}\|^2 - \|\mathbf{y} - \hat{\mu}_k\|^2)}\right]$$

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- As with James-Stein shrinkage, gap tends to be larger when signal-to-noise ratio $\|\mu\|/\sigma$ or sample size are small

In past statistical work, all the focus has been on selecting the "best" single model from a class of models. We may need to shift our thinking to the possibility of forming combinations of models... — Leo Breiman (1996)

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- Connect to other ensemble methods like random forests (randomization + model selection)

Chen, K., & Tan, Error Reduction from Stacked Regressions (2023)

Available at klusowski.princeton.edu

